

Robust Standard Errors in Small Samples

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Description

This package implements the small-sample degrees of freedom adjustment discussed in Imbens and Kolesár [2016]. The implementation can handle models with fixed effects, and cases where the number of observations or clusters is large ¹.

```
library(dfadjust)
```

To give some examples, let us construct an artificial dataset with 11 clusters

```
set.seed(7)
d1 <- data.frame(y = rnorm(1000), x1 = c(rep(1, 3), rep(0,
  997)), x2 = c(rep(1, 150), rep(0, 850)), x3 = rnorm(1000),
  cl = as.factor(c(rep(1:10, each = 50), rep(11, 50))))
```

Let us first run a regression of y on x_1 . This is a case where in spite of moderate data size, the effective number of observations is small since there are only three treated units:

```
r1 <- lm(y ~ x1, data = d1)
## No clustering
dfadjustSE(r1)
#>
#> Coefficients:
#>           Estimate HC1 se HC2 se Adj. se      df p-value
#> (Intercept)  0.00266 0.0311  0.031  0.0311 996.00  0.932
#> x1          0.12940 0.8892  1.088  2.3743   2.01  0.957
```

Now consider a cluster-robust regression of y on x_2 . There are only 3 treated clusters, so the effective number of observations is again small:

¹We thank Ulrich Müller for suggesting to us the lemma below

```

r1 <- lm(y ~ x2, data = d1)
# Default Imbens-Kolesár method
dfadjustSE(r1, clustervar = d1$c1)
#>
#> Coefficients:
#>           Estimate HC1 se HC2 se Adj. se   df p-value
#> (Intercept) -0.0236 0.0135 0.0169 0.0222 4.94 0.288
#> x2           0.1778 0.0530 0.0621 0.1157 2.43 0.124
# Bell-McCaffrey method
dfadjustSE(r1, clustervar = d1$c1, IK = FALSE)
#>
#> Coefficients:
#>           Estimate HC1 se HC2 se Adj. se   df p-value
#> (Intercept) -0.0236 0.0135 0.0169 0.0316 2.42 0.4547
#> x2           0.1778 0.0530 0.0621 0.1076 2.70 0.0983

```

Now, let us run a regression of y on x_3 , with fixed effects. Since we're only interested in x_3 , we specify that we only want inference on the second element:

```

r1 <- lm(y ~ x3 + c1, data = d1)
dfadjustSE(r1, clustervar = d1$c1, ell = c(0, 1, rep(0,
  r1$rank - 2)))
#>
#> Coefficients:
#>           Estimate HC1 se HC2 se Adj. se   df p-value
#> Estimate    0.0261 0.0463 0.0595 0.0928 3.23 0.778
dfadjustSE(r1, clustervar = d1$c1, ell = c(0, 1, rep(0,
  r1$rank - 2)), IK = FALSE)
#>
#> Coefficients:
#>           Estimate HC1 se HC2 se Adj. se   df p-value
#> Estimate    0.0261 0.0463 0.0595 0.0928 3.23 0.778

```

Finally, an example in which the clusters are large. We have 500,000 observations:

```

d2 <- do.call("rbind", replicate(500, d1, simplify = FALSE))
d2$y <- rnorm(length(d2$y))
r2 <- lm(y ~ x2, data = d2)
summary(r2)
#>
#> Call:
#> lm(formula = y ~ x2, data = d2)
#>
#> Residuals:
#>   Min       1Q   Median       3Q      Max
#> -5.073 -0.675  0.000  0.675  4.789
#>
#> Coefficients:
#>           Estimate Std. Error t value Pr(>|t|)

```

```

#> (Intercept) -0.000991  0.001535  -0.65    0.52
#> x2          -0.003590  0.003963  -0.91    0.37
#>
#> Residual standard error: 1 on 499998 degrees of freedom
#> Multiple R-squared:  1.64e-06,   Adjusted R-squared:  -3.59e-07
#> F-statistic: 0.821 on 1 and 5e+05 DF,  p-value: 0.365
# Default Imbens-Kolesár method
dfadjustSE(r2, clustervar = d2$c1)
#>
#> Coefficients:
#>              Estimate HC1 se  HC2 se Adj. se   df p-value
#> (Intercept) -0.000991 0.00133 0.00205 0.00261 5.50   0.704
#> x2          -0.003590 0.00483 0.00376 0.00554 3.64   0.517
# Bell-McCaffrey method
dfadjustSE(r2, clustervar = d2$c1, IK = FALSE)
#>
#> Coefficients:
#>              Estimate HC1 se  HC2 se Adj. se   df p-value
#> (Intercept) -0.000991 0.00133 0.00205 0.00267 5.10   0.710
#> x2          -0.003590 0.00483 0.00376 0.00554 3.64   0.517

```

Methods

This section describes the implementation of the Imbens and Kolesár [2016] and Bell and McCaffrey [2002] degrees of freedom adjustments.

There are S clusters, and we observe n_s observations in cluster s , for a total of $n = \sum_{s=1}^S n_s$ observations. We handle the case with independent observations by letting each observation be in its own cluster, with $S = n$. Consider the linear regression of a scalar outcome Y_i onto a p -vector of regressors X_i ,

$$Y_i = X_i' \beta + u_i, \quad E[u_i | X_i] = 0.$$

We're interested in inference on $\ell' \beta$ for some fixed vector $\ell \in \mathbb{R}^p$. Let X , u , and Y denote the design matrix, and error and outcome vectors, respectively. For any $n \times k$ matrix M , let M_s denote the $n_s \times k$ block corresponding to cluster s , so that, for instance, Y_s corresponds to the outcome vector in cluster s . For a positive semi-definite matrix M , let $M^{1/2}$ be a matrix satisfying $M^{1/2'} M^{1/2} = M$, such as its symmetric square root or its Cholesky decomposition.

Assume that

$$E[u_s u_s' | X] = \Omega_s, \quad \text{and} \quad E[u_s u_t' | X] = 0 \quad \text{if } s \neq t.$$

Denote the conditional variance matrix of u by Ω , so that Ω_s is the block of Ω corresponding to cluster s . We estimate $\ell' \beta$ using OLS. In R, the OLS estimator is computed via a QR decomposition, $X = QR$, where $Q'Q = I$ and R is upper-triangular, so we can write the estimator as

$$\ell' \hat{\beta} = \ell' \left(\sum_s X_s' X_s \right)^{-1} \sum_s X_s Y_s = \tilde{\ell}' \sum_s Q_s' Y_s, \quad \tilde{\ell} = R^{-1} \ell.$$

It has variance

$$V := \text{var}(\ell' \hat{\beta} \mid X) = \ell' (X'X)^{-1} \sum_s X'_s \Omega_s X_s (X'X)^{-1} \ell = \tilde{\ell}' \sum_s Q'_s \Omega_s Q_s \tilde{\ell}.$$

Variance estimate

We estimate V using a variance estimator that generalizes the HC2 variance estimator to clustering. Relative to the LZ2 estimator described in Imbens and Kolesár [2016], we use a slight modification that allows for fixed effects:

$$\hat{V} = \ell' (X'X)^{-1} \sum_s X'_s A_s \hat{u}_s \hat{u}'_s A'_s X_s (X'X)^{-1} \ell = \ell' R^{-1} \sum_s Q'_s A_s \hat{u}_s \hat{u}'_s A'_s Q_s R^{-1} \ell = \sum_{s=1}^S (\hat{u}'_s a_s)^2,$$

where

$$\hat{u}_s := Y_s - X_s \hat{\beta} = u_s - Q_s Q' u, \quad a_s = A'_s Q_s \tilde{\ell},$$

and the matrix A_s is given by the symmetric square root of the inverse of $I - Q_s Q'_s$, or else its pseudo-inverse if it is singular, as is the case, for example, if X contains fixed effects. We do not need to insist on $I - Q_s Q'_s$ to be invertible, since, using the identity

$$\hat{V} = u \sum_s (I - Q Q')'_s a_s a'_s (I - Q Q')_s u,$$

one can verify by simple algebra that a sufficient condition for \hat{V} to be unbiased under homoskedasticity is that $Q'_s A_s (I - Q_s Q'_s) A_s Q_s = Q'_s Q_s$ (see, for example, Pustejovsky and Tipton [2018], for details).

If the observations are independent, the vector of leverages $(Q'_1 Q_1, \dots, Q'_n Q_n)$ can be computed directly using the `stats::hatvalues` function. In this case, use this function to compute $A_i = 1/\sqrt{1 - Q'_i Q_i}$ directly, and we then compute $a_i = A_i Q'_i \tilde{\ell}$ using vector operations. For the case with clustering, computing the spectral decomposition of $I - Q_s Q'_s$ can be expensive or even infeasible if the cluster size n_s is large. We therefore use the following result, suggested to us by Ulrich Müller, allows us to compute a_s by computing a spectral decomposition of a $p \times p$ matrix.

- Let $Q'_s Q_s = \sum_{i=1}^p \lambda_{is} r_{is} r'_{is}$ be the spectral decomposition of $Q'_s Q_s$. Then $A_s = \sum_{i: \lambda_i \neq 1} (1 - \lambda_i)^{-1/2} Q_s r_{is} r'_{is} Q'_s$, satisfies $A_s (I - Q_s Q'_s) A_s = I$.

This follows from the fact that $I - Q_s Q'_s$ has eigenvalues $1 - \lambda_{is}$ and eigenvectors $Q_s r_{is}$, and hence its pseudoinverse is $\sum_{i: \lambda_i \neq 1} (1 - \lambda_i)^{-1} Q_s r_{is} r'_{is} Q'_s$.

Using the lemma, we can compute a_s efficiently as:

$$a_s = \sum_{i: \lambda_i \neq 1} (1 - \lambda_i)^{-1/2} Q_s r_{is} r'_{is} Q'_s Q_s \tilde{\ell} = Q_s D_s \tilde{\ell}, \quad D_s = \sum_{i: \lambda_i \neq 1} \lambda_i (1 - \lambda_i)^{-1/2} r_{is} r'_{is}.$$

Degrees of freedom correction

Let G be an $n \times S$ matrix with columns $(I - Q Q')'_s a_s$. Then the Bell and McCaffrey [2002] adjustment sets the degrees of freedom to

$$f_{\text{BM}} = \frac{\text{tr}(G'G)^2}{\text{tr}((G'G)^2)}.$$

Since $(G'G)_{st} = a'_s(I - QQ')_s(I - QQ')'_t a_t = a_s(\mathbb{1}\{s = t\} - Q_s Q'_t) a_t$, the matrix $G'G$ can be efficiently computed as

$$G'G = \text{diag}(a'_s a_s) - BB' \quad B_{sk} = a'_s Q_{sk}.$$

Note that B is an $S \times p$ matrix, so that computing the degrees of freedom adjustment only involves $p \times p$ matrices:

$$f_{\text{BM}} = \frac{(\sum_s a'_s a_s - \sum_{s,k} B_{sk}^2)^2}{\sum_s (a'_s a_s)^2 - 2 \sum_{s,k} (a'_s a_s) B_{sk}^2 + \sum_{s,t} (B'_s B_t)^2}.$$

If the observations are independent, we compute B directly as $B \leftarrow a * Q$, and since a_i is a scalar, we have

$$f_{\text{BM}} = \frac{(\sum_i a_i^2 - \sum_{s,k} B_{sk}^2)^2}{\sum_i a_i^4 - 2 \sum_i a_i^2 B'_i B_i + \sum_{i,j} (B'_i B_j)^2}.$$

The Imbens and Kolesár [2016] degrees of freedom adjustment instead sets

$$f_{\text{IK}} = \frac{\text{tr}(G' \hat{\Omega} G)^2}{\text{tr}((G' \hat{\Omega} G)^2)},$$

where $\hat{\Omega}$ is an estimate of the Moulton [1986] model of the covariance matrix, under which $\Omega_s = \sigma_\epsilon^2 I_{n_s} + \rho l_{n_s} l'_{n_s}$. Using simple algebra, one can show that in this case,

$$G' \Omega G = \sigma_\epsilon^2 \text{diag}(a'_s a_s) - \sigma_\epsilon^2 BB' + \rho(D - BF')(D - BF')',$$

where

$$F_{sk} = l'_{n_s} Q_{sk}, \quad D = \text{diag}(a'_s l_{n_s})$$

which can again be computed even if the clusters are large. The estimate $\hat{\Omega}$ replaces σ_ϵ^2 and ρ with analog estimates.

References

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- James E. Pustejovsky and Elizabeth Tipton. Small-sample methods for cluster-robust variance estimation and hypothesis testing in fixed effects models. *Journal of Business & Economic Statistics*, 36(4):672–683, October 2018. doi: 10.1080/07350015.2016.1247004.